# Recent Advances in High-Dimensional Robust Statistics

Ilias Diakonikolas (UW Madison) ICML 2020 Tutorial July 2020 PART II: BASIC ALGORITHMIC TECHNIQUES

# OUTLINE

#### Part II

- Basics: Sample Complexity of Robust Estimation
- Certificate of Robustness
- Recursive Dimension Halving
- Iterative Filtering

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### BASICS OF HIGH-DIMENSIONAL MEAN ESTIMATION

**Fact**: Let  $X_1, \ldots, X_N$  be IID samples from  $\mathcal{N}(\mu, I)$ . The empirical estimator  $\hat{\mu}$  satisfies  $\|\hat{\mu} - \mu\|_2 = O(\sqrt{d/N})$  with probability at least 9/10. Moreover, this rate is optimal for any estimator.

#### **Proof:**

By definition, 
$$\hat{\mu} = (1/N) \sum_{i=1}^{N} X_i$$
, where  $X_i \sim \mathcal{N}(\mu, I)$ .  
Then,  
 $\hat{\mu} \sim \mathcal{N}(\mu, (1/N)I)$ .

We have

Therefore,

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$$\begin{split} \|\widehat{\mu} - \mu\|_{2}^{2} &= \sum_{j=1}^{d} \mathbf{E} \left[ (\widehat{\mu}_{j} - \mu_{j})^{2} \right] = \sum_{j=1}^{d} \mathbf{Var} \left[ \widehat{\mu}_{j} \right] \ = d/N \\ \|\widehat{\mu} - \mu\|_{2}^{2} &\leq \mathbf{E} [\|\widehat{\mu} - \mu\|_{2}^{2}]^{1/2} = \sqrt{\frac{d}{N}} \end{split}$$

For lower bound, apply Assouad's lemma for  $\mathcal{M} = \{\mu_j = \pm \sqrt{N}/d, j \in [d]\}$ 

### INFORMATION-THEORETIC LIMITS ON ROBUST ESTIMATION

**Proposition**: Any robust mean estimator for  $\mathcal{N}(\mu, 1)$  has error  $\Omega(\epsilon)$ , even in Huber's model.

**Claim**: Let  $P_1$ ,  $P_2$  be such that  $d_{TV}(P_1, P_2) = \epsilon/(1 - \epsilon)$ . There exist noise distributions  $B_1$ ,  $B_2$  such that  $(1 - \epsilon)P_1 + \epsilon B_1 = (1 - \epsilon)P_2 + \epsilon B_2$ .

- Use  $d_{\mathrm{TV}}(\mathcal{N}(\mu_1, 1), \mathcal{N}(\mu_2, 1)) \le |\mu_1 \mu_2|/2$
- Same argument gives:
  - For sub-gaussian distributions:  $\Omega(\epsilon \sqrt{\log(1/\epsilon)})$
  - For bounded variance distributions:  $\Omega(\sqrt{\epsilon})$

# SAMPLE EFFICIENT ROBUST MEAN ESTIMATION (I)

**Proposition**: There is an algorithm that uses  $N = O(d/\epsilon^2) \epsilon$ - corrupted samples from  $\mathcal{N}(\mu, I)$  and outputs  $\tilde{\mu} \in \mathbb{R}^d$  that with probability at least 9/10 satisfies  $\|\tilde{\mu} - \mu\|_2 = O(\epsilon)$ .

**Main Idea**: To robustly learn the mean of  $\mathcal{N}(\mu, I)$ , it suffices to learn the mean of *all* its 1-dimensional projections (cf. Tukey median).

**Basic Fact**: 
$$\|x\|_2 = \max_{v: \|v\|_2 = 1} |v \cdot x|$$

**Claim 1**: Suppose we can find  $\{\widehat{\mu}_v\}_v$  s.t. for all  $v \in \mathbb{R}^d$  with  $||v||_2 = 1$  we have  $|\widehat{\mu}_v - \mu \cdot v| \leq \delta$ . Then, we can estimate  $\mu$  within error  $2\delta$ .

#### **Proof:**

Consider *infinite size* LP: Find  $x \in \mathbb{R}^d$  such that for all unit  $v \in \mathbb{R}^d$ :  $|\hat{\mu}_v - v \cdot x| \leq \delta$ .

Let  $x^*$  be any feasible solution. Then

$$\|x^* - \mu\|_2 = \max_{v: \|v\|_2 = 1} |v \cdot x^* - v \cdot \mu| \le \max_{v: \|v\|_2 = 1} |v \cdot x^* - \widehat{\mu}_v| + \max_{v: \|v\|_2 = 1} |v \cdot \mu - \widehat{\mu}_v| \le 2\delta .$$

### SAMPLE EFFICIENT ROBUST MEAN ESTIMATION (II)

**Main Idea**: To robustly learn the mean of  $\mathcal{N}(\mu, I)$ , it suffices to learn the mean of "all" its 1-dimensional projections.

**Claim 2**: Suffices to consider a  $\gamma$ -net *C* over all directions, where  $\gamma$  is a small positive constant. **Proof:** 

This gives *finite* LP:

Find  $x \in \mathbb{R}^d$  such that for all  $v \in C$ , we have  $|\widehat{\mu}_v - v \cdot x| \leq \delta$ .

Let  $x^*$  be any feasible solution. Let  $u \in C$  such that  $||u - \frac{\mu - x^*}{||\mu - x^*||_2}||_2 \leq \gamma$ . Then

$$\begin{split} \|x^* - \mu\|_2 &= \left| \left( \left( \frac{\mu - x^*}{\|\mu - x^*\|_2} - u \right) + u \right) \cdot (x^* - \mu) \right| \le \gamma \|x^* - \mu\|_2 + 2\delta \\ \|x^* - \mu\|_2 \le \frac{2\delta}{1 - \gamma} \ . \end{split}$$

or

# SAMPLE EFFICIENT ROBUST MEAN ESTIMATION (III)

**Main Idea**: To robustly learn the mean of  $\mathcal{N}(\mu, I)$ , it suffices to learn the mean of "all" its 1-dimensional projections.

So, for  $\gamma = 1/2$ , any feasible solution to the LP has  $||x^* - \mu||_2 \le 4\delta$ .

**Sample Complexity**: Note that the median satisfies  $\delta = O(\epsilon)$  with probability at least  $1 - \tau$  using  $O((1/\epsilon^2) \log(1/\tau))$  samples.

We need union bound over all  $v \in C$ . Since  $|C| = (1/\gamma)^{O(d)} = 2^{O(d)}$ , for  $\tau = 1/(10|C|)$  algorithm works with probability at least 9/10. Thus, sample complexity will be  $N = O(d/\epsilon^2)$ .

**Runtime**:  $poly(N, 2^d)$ .

# OUTLINE

#### Part II

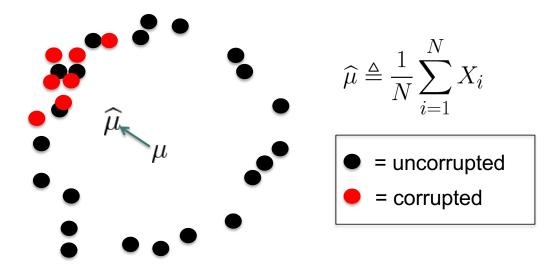
- Basics: Sample Complexity of Robust Estimation
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### CERTIFICATE OF ROBUSTNESS FOR EMPIRICAL ESTIMATOR

**Idea #1 [DKKLMS'16, LRV'16]**: If the empirical covariance is "close to what it should be", then the empirical mean works.

# CERTIFICATE FOR EMPIRICAL MEAN

Detect when the empirical estimator may be compromised



There is no direction of large empirical variance

**Key Lemma:** Let  $X_1, X_2, ..., X_N$  be an  $\epsilon$ -corrupted set of samples from  $\mathcal{N}(\mu, I)$  and  $N = \Omega(d/\epsilon^2)$ , then for

(1) 
$$\widehat{\mu} \triangleq \frac{1}{N} \sum_{i=1}^{N} X_i$$
 (2)  $\widehat{\Sigma} \triangleq \frac{1}{N} \sum_{i=1}^{N} (X_i - \widehat{\mu}) (X_i - \widehat{\mu})^T$ 

with high probability we have:

$$\|\widehat{\Sigma}\|_2 \le 1 + O(\epsilon \log(1/\epsilon)) \longrightarrow \|\widehat{\mu} - \mu\|_2 \le O(\epsilon \sqrt{\log(1/\epsilon)})$$

in strong contamination model.

Key Lemma: Let  $X_1, X_2, ..., X_N$  be an  $\epsilon$ -corrupted set of samples from  $\mathcal{N}(\mu, I)$ and  $N = \Omega(d/\epsilon^2)$ , then for (1)  $\widehat{\mu} \triangleq \frac{1}{N} \sum_{i=1}^{N} X_i$  (2)  $\widehat{\Sigma} \triangleq \frac{1}{N} \sum_{i=1}^{N} (X_i - \widehat{\mu})(X_i - \widehat{\mu})^T$ with high probability we have:  $\|\widehat{\Sigma}\|_2 \le 1 + \delta \longrightarrow \|\widehat{\mu} - \mu\|_2 \le O(\sqrt{\delta\epsilon} + \epsilon \sqrt{\log(1/\epsilon)})$ in strong contamination model.

Idea #2 [DKKLMS'16]: Removing any  $\epsilon$ -fraction of good points does not move the empirical mean and covariance by much.

### **REMARKS ON KEY LEMMA**

- Statement applies for spherical distributions with sub-Gaussian tails.
- Essentially same argument goes through if covariance is *approximately* known.
- Argument extends for distributions with known covariance and weaker concentration.

If *D* is isotropic with *sub-exponential* tails:

$$\|\widehat{\Sigma}\|_2 \le 1 + \delta \longrightarrow \|\widehat{\mu} - \mu\|_2 \le O(\sqrt{\delta\epsilon} + \epsilon \log(1/\epsilon))$$

If *D* satisfies 
$$\Sigma \leq I$$
:  
 $\|\widehat{\Sigma}\|_2 \leq 1 + \delta \implies \|\widehat{\mu} - \mu\|_2 \leq O(\sqrt{\delta\epsilon} + \sqrt{\epsilon}).$ 

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Idea #3 [LRV'16]: Additive corruptions can move the covariance in *some* directions, but *not in all* directions simultaneously.

### **RECURSIVE DIMENSION-HALVING [LRV'16]**

#### **LRV Procedure:**

**Step #1:** Find large subspace where "standard" estimator works. **Step #2:** Recurse on complement.

Combine Results.

Can reduce dimension by factor of 2 in each recursive step.

# FINDING A GOOD SUBSPACE (I)

"Good subspace G" = one where the empirical mean works

By Key Lemma, sufficient condition is:

Projection of empirical covariance on **G** has no large eigenvalues.

• Also want **G** to be "high-dimensional".

Question: How do we find such a subspace?

### FINDING A GOOD SUBSPACE (II)

**Good Subspace Lemma:** Let  $X_1, X_2, ..., X_N$  be an *additively*  $\epsilon$ -corrupted set of  $N = \Omega(d \log d/\epsilon^2)$  samples from  $\mathcal{N}(\mu, I)$ . *After naïve pruning*, we have that  $\lambda_{d/2}(\widehat{\Sigma}) \leq 1 + O(\epsilon)$ 

**Corollary:** Let *W* be the span of the bottom d/2 eigenvalues of  $\widehat{\Sigma}$ . Then *W* is a good subspace.

# **RECURSIVE DIMENSION-HALVING ALGORITHM [LRV'16]**

Algorithm works as follows:

- Remove gross outliers (e.g., naïve pruning).
- Let W, V be the span of bottom d/2 and upper d/2 eigenvalues of  $\widehat{\Sigma}$  respectively.
- Use empirical mean on *W*.
- Recurse on *V* (If the dimension is one, use median).

#### Error Analysis:

 $O(\log d)$  levels of the recursion  $\longrightarrow$  final error of  $O(\epsilon \sqrt{\log d})$ 

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Idea #4 [DKKLMS'16]: Iteratively "remove outliers" in order to "fix" the empirical covariance.

# **ITERATIVE FILTERING [DKKLMS'16]**

#### **Iterative Two-Step Procedure:**

**Step #1:** Test certificate of robustness of "standard" estimator

Step #2: If certificate is violated, detect and remove outliers

Iterate on "cleaner" dataset.

General recipe that works in general settings.

Let's see how this works for robust mean estimation.

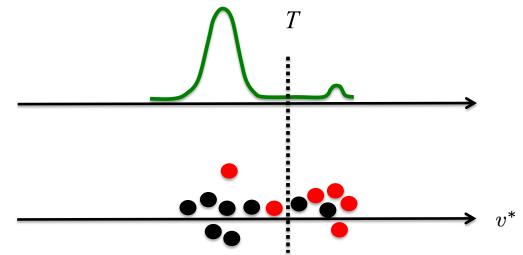
### FILTERING SUBROUTINE

Either output empirical mean, or remove many outliers.

Filtering Approach: Suppose that:

$$\|\widehat{\Sigma}\|_2 \ge 1 + \Omega(\epsilon \log(1/\epsilon))$$

Let  $v^*$  be the direction of maximum variance.



### FILTERING SUBROUTINE

Either output empirical mean, or remove many outliers.

Filtering Approach: Suppose that:

$$\|\widehat{\Sigma}\|_2 \ge 1 + \Omega(\epsilon \log(1/\epsilon))$$

Let  $v^*$  be the direction of maximum variance.

- Project all the points on the direction of  $v^*$ .
- Find a threshold *T* such that

$$\mathbf{Pr}_{X \sim_U S}[|v^* \cdot X - \text{median}(\{v^* \cdot x, x \in S\})| > T+1] \ge 8 \cdot e^{-T^2/2}$$

• Throw away all points *x* such that

$$v^* \cdot x - \text{median}(\{v^* \cdot x, x \in S\})| > T + 1$$

• Iterate on new dataset.

### FILTERING SUBROUTINE: ANALYSIS SKETCH

Either output empirical mean, or remove many outliers.

Filtering Approach: Suppose that:

$$\|\widehat{\Sigma}\|_2 \ge 1 + \Omega(\epsilon \log(1/\epsilon))$$

Claim: In each iteration, we remove more outliers than inliers.

After a bounded number of iterations, we stop removing points.

Eventually the empirical mean works

**Runtime:**  $\tilde{O}(Nd^2)$ 

### FILTERING PSEUDO-CODE

**Input**:  $\epsilon$ -corrupted set *S* from  $\mathcal{N}(\mu, I)$  **Output**: Set  $S' \subseteq S$  that is  $\epsilon'$ -corrupted, for some  $\epsilon' < \epsilon$ OR robust estimate of the unknown mean  $\mu$ 

- **1.** Let  $\hat{\mu}_S, \hat{\Sigma}_S$  be the empirical mean and covariance of the set *S*.
- 2. If  $\|\widehat{\Sigma}_S\|_2 \le 1 + C\epsilon \log(1/\epsilon)$ , for an appropriate constant C > 0: Output  $\widehat{\mu}_S$
- **3.** Otherwise, let  $(\lambda^*, v^*)$  be the top eigenvalue-eigenvector pair of  $\widehat{\Sigma}_{S^*}$ .
- **4.** Find T > 0 such that

 $\mathbf{Pr}_{X \sim_U S}[|v^* \cdot X - \text{median}(\{v^* \cdot x, x \in S\})| > T+1] \ge 8 \cdot e^{-T^2/2}.$ 

5. Return

$$S' = \{x \in S : |v^* \cdot x - \text{median}(\{v^* \cdot x, x \in S\})| \le T + 1\}.$$

### **REMARKS ON FILTERING METHOD(S)**

- For known covariance sub-Gaussian case, filter relied on violation of concentration.
- This extends to weaker concentration, as long as covariance is (approximately) known.
- For example, for *sub-exponential* concentration, filter would be:

Find T > 0 such that  $\mathbf{Pr}_{X \sim_U S}[|v^* \cdot (X - \widehat{\mu})| > T] \ge 8 \cdot e^{-T}$ .

• For the bounded covariance setting, randomized filtering / down-weighting.

Remove point x with probability proportional to  $(v^* \cdot (x - \widehat{\mu}))^2$  .

• Analogue of Claim 1: Remove more outliers than inliers in expectation.

# SUMMARY: ROBUST MEAN ESTIMATION VIA FILTERING

#### Certificate for Robustness:

"Spectral norm of empirical covariance is *close* to what it should be."

#### **Exploiting the Certificate:**

- Check if certificate is satisfied.
- If violated, find "subspace" where behavior of outliers different than behavior of inliers.
- Use it to detect and remove outliers.
- Iterate on "cleaner" dataset.