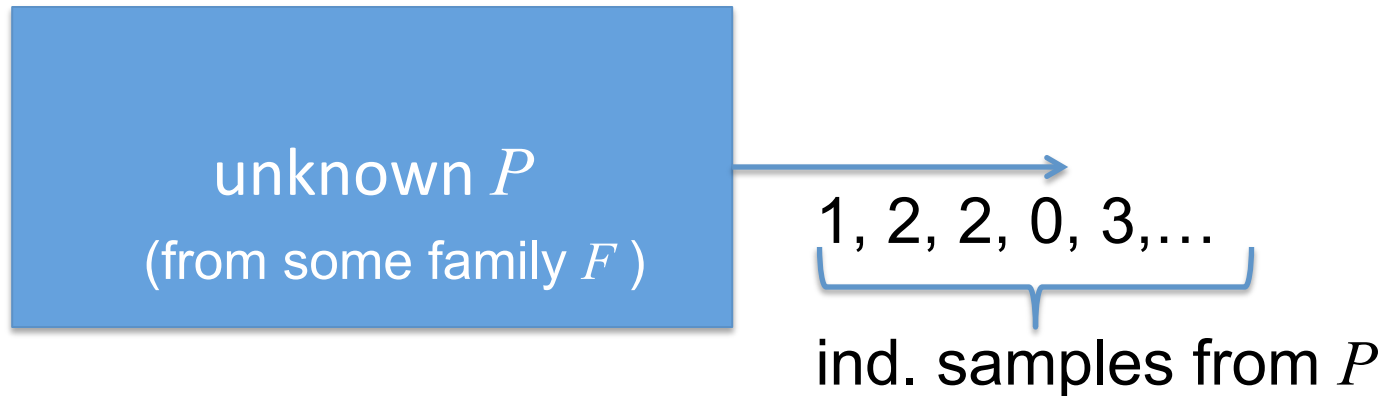


Beyond Berry-Esseen:
Structure and Learning
Sums of Random Variables

Constantinos Daskalakis
EECS, MIT

Distribution Learning Problem

- **Input:** - Sample access to distribution over $\{0,1,\dots,n\}$



– $\varepsilon > 0$

- **Goal:**

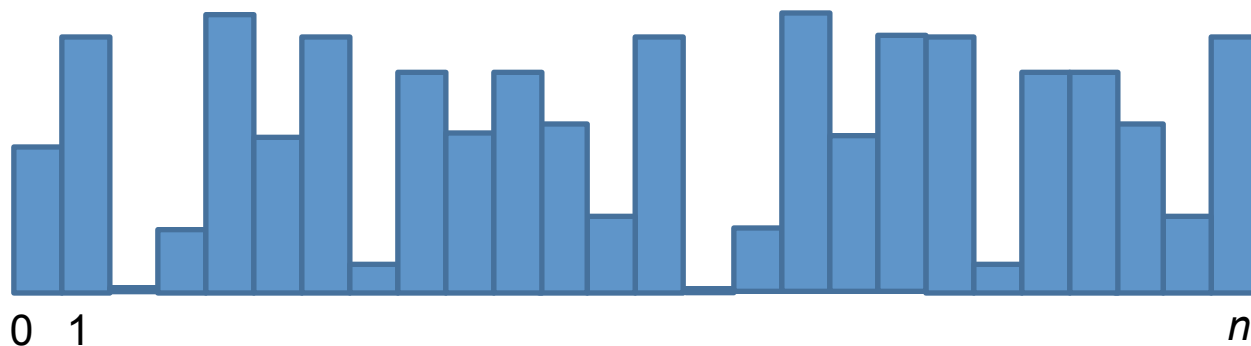
$$\frac{1}{2} \sum_{i=0}^n |P(i) - Q(i)|$$

- Find some Q s.t. $d_{TV}(P, Q) \leq \varepsilon$
- **(proper learn)** Find $Q \in F$ s.t. $d_{TV}(P, Q) \leq \varepsilon$
- **Minimize number of samples, computation time**

Distribution Learning Problem (cont'd)

- **Refresher:**

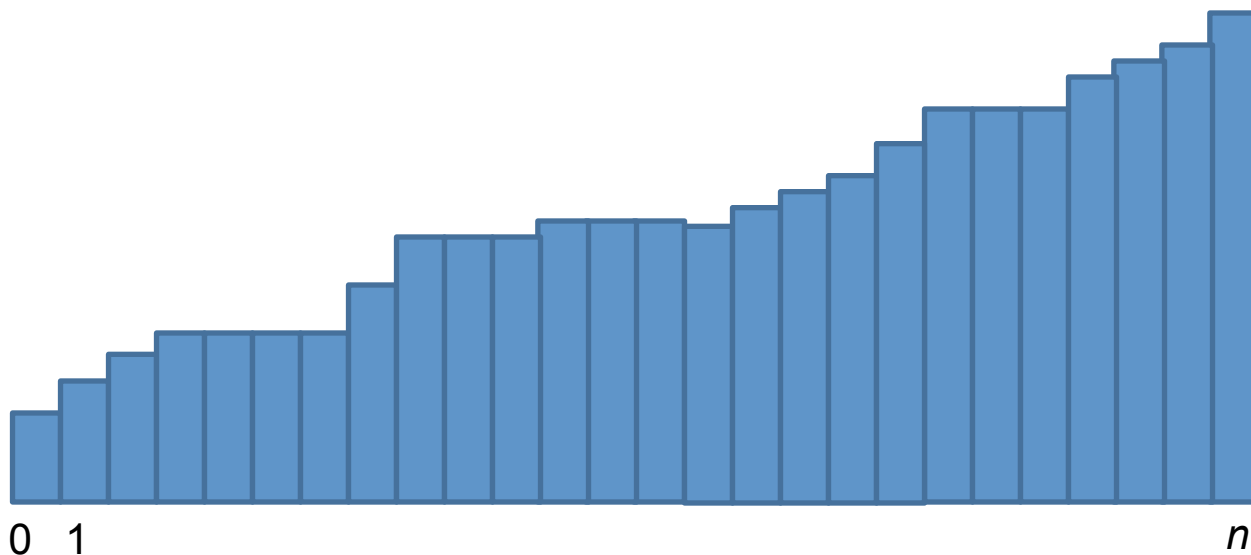
- Arbitrary distribution over $\{0, \dots, n\}$ requires time and sample complexity of $\Theta(n/\epsilon^2)$ (folklore)



Distribution Learning Problem (cont'd)

- **Refresher:**

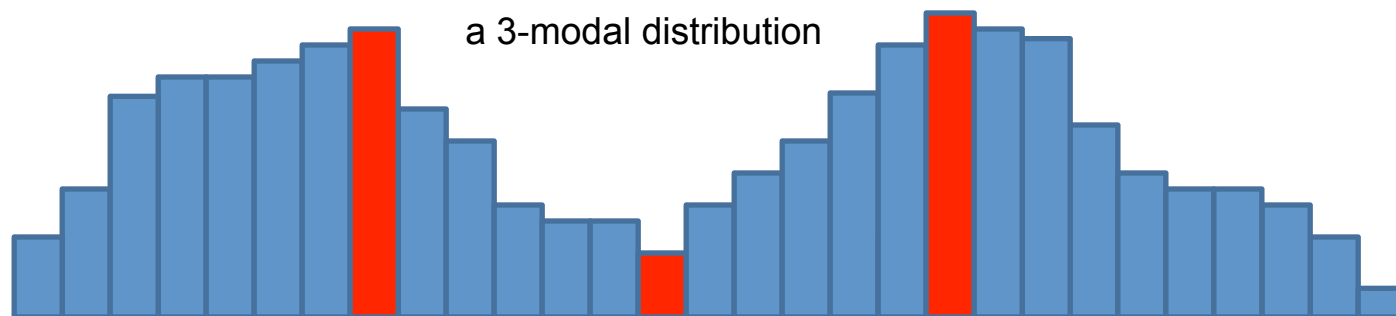
- Arbitrary distribution over $\{0, \dots, n\}$ requires time and sample complexity of $\Theta(n/\epsilon^2)$ (folklore)
- Monotone distributions over $\{0, \dots, n\}$ require time and sample complexity of $\Theta(\log n/\epsilon^3)$ [Birgé 1987]



Distribution Learning Problem (cont'd)

- **Refresher:**

- Arbitrary distribution over $\{0, \dots, n\}$ requires time and sample complexity of $\Theta(n/\epsilon^2)$ (folklore)
- Monotone distributions over $\{0, \dots, n\}$ require time and sample complexity of $\Theta(\log n/\epsilon^3)$ [Birgé 1987]
- k -modal distributions over $\{0, \dots, n\}$ can be learned from $O\left(\frac{k \log n}{\epsilon^3} + \frac{k^3 \log k/\epsilon}{\epsilon^3}\right)$ samples in time $\text{poly}(k \log n/\epsilon)$ [D-Diakonikolas-Servedio 2012]



Distribution Learning Problem (cont'd)

- **Refresher:**

- Arbitrary distribution over $\{0, \dots, n\}$ requires time and sample complexity of $\Theta(n/\epsilon^2)$ (folklore)
- Monotone distributions over $\{0, \dots, n\}$ require time and sample complexity of $\Theta(\log n/\epsilon^3)$ [Birgé 1987]
- k -modal distributions over $\{0, \dots, n\}$ can be learned from $O\left(\frac{k \log n}{\epsilon^3} + \frac{k^3 \log k/\epsilon}{\epsilon^3}\right)$ samples in time $\text{poly}(k \log n/\epsilon)$ [D-Diakonikolas-Servedio 2012]
- Log-concave distributions $P(i)^2 > P(i-1) P(i+1)$ can be learned from $\tilde{O}(1/\epsilon^{2.5})$ samples [Chan-Diakonikolas-Servedio-Sun 2014]

Focus of This Talk: PBDs and SIIRVs

- **Def 1:** A Poisson Binomial Distribution (PBD) is
 - the distribution of the sum $X = \sum X_i$ of n independent r.v.'s $X_i \in \{0,1\}$
 - support: $\{0,1,\dots,n\}$

Sharp structural results

- **Def 2:** A k -SIIRV is
 - the distribution of the sum $X = \sum X_i$ of n independent r.v.'s $X_i \in \{0,\dots,k-1\}$
 - support: $\{0,1,\dots,n \cdot (k-1)\}$

Learning from $\Theta(1/\epsilon^2)$ /
respectively $\text{poly}(k/\epsilon)$ samples

- **Objectives:** *Structure* and *Learning*

Menu

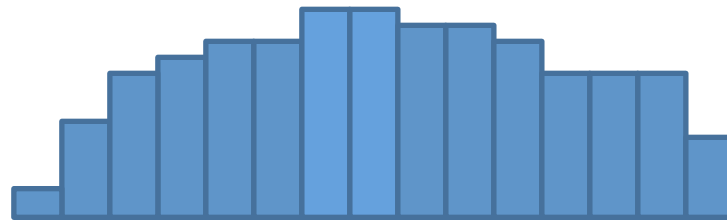
- *Refresher*
- *Objectives for this talk*
- PBD Structure and Learning
- k -SIIRV Structure and Learning

Menu

- *Refresher*
- *Objectives for this talk*
- **PBD Structure and Learning**
- *k*-SIIRV Structure and Learning

Structure of PBDs

- Unimodal; in fact log-concave distributions
 - so can be (non-properly) learned from $\tilde{O}(1/\epsilon^{2.5})$ samples



- [Berry 1941, Esseen 1942]: If X_1, \dots, X_n are independent and bounded then

$$d_K \left(\sum_i X_i, \mathcal{N}(\mu, \sigma^2) \right) \leq C \frac{\sum_i \mathbb{E}[|X_i|^3]}{\sigma^3}$$

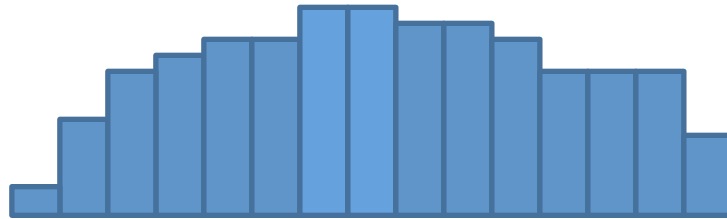
$$d_K(P, Q) = \max_i |P(\leq i) - Q(\leq i)|$$

$$\mu = \mathbb{E} \left[\sum_i X_i \right], \sigma^2 = \text{Var} \left[\sum_i X_i \right]$$

$$[\text{Esseen 1956}] \quad 0.4097 \leq C \leq 0.5600 \quad [\text{Shevtsova 2010}]$$

Structure of PBDs

- Unimodal; in fact log-concave distributions
 - so can be (non-properly) learned from $\tilde{O}(1/\epsilon^{2.5})$ samples



- [Berry 1941, Esseen 1942]: If X_1, \dots, X_n are independent and bounded then

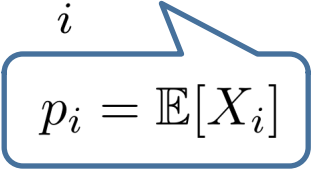
$$d_K \left(\sum_i X_i, \mathcal{N}(\mu, \sigma^2) \right) \leq C \frac{\sum_i \mathbb{E}[|X_i|^3]}{\sigma^3}$$

- specializing to PBDs: $d_K \left(\sum_i X_i, \mathcal{N}(\mu, \sigma^2) \right) \leq C \frac{\mu}{\sigma^3}$
- e.g. $d_K(B(n, p), \mathcal{N}(np, np(1-p))) \leq C \frac{1}{\sqrt{np}(1-p)^{1.5}}$

quality of bound decays with $n \cdot p$ – poor if, eg, $p = 1/n$

Structure of PBDs

- [Berry 1941, Esseen 1942]: $d_K \left(\sum_i X_i, N(\mu, \sigma^2) \right) \leq C \frac{\mu}{\sigma^3}$
- e.g. $d_K(B(n, p), \mathcal{N}(np, np(1-p))) \leq C \frac{1}{\sqrt{np}(1-p)^{1.5}}$
- [Le Cam 1960]: $d_{TV} \left(\sum_i X_i, \text{Poisson}(\mu) \right) \leq \sum_i p_i^2$
- e.g. $d_{TV}(B(n, p), \text{Poisson}(np)) \leq np^2$
- good when, e.g., $p = 1/n$


$$p_i = \mathbb{E}[X_i]$$

Structure of PBDs

- [Berry 1941, Esseen 1942]: $d_K \left(\sum_i X_i, N(\mu, \sigma^2) \right) \leq C \frac{\mu}{\sigma^3}$
- [Le Cam 1960]: $d_{TV} \left(\sum_i X_i, \text{Poisson}(\mu) \right) \leq \sum_i p_i^2$
- [Chen-Goldstein-Shao 2011]: $d_{TV} \left(\sum_i X_i, Z(\mu, \sigma^2) \right) \leq \frac{O(1)}{\sigma}$

rounded $\mathcal{N}(\mu, \sigma^2)$

Structure of PBDs

- [Berry 1941, Esseen 1942]: $d_K \left(\sum_i X_i, N(\mu, \sigma^2) \right) \leq C \frac{\mu}{\sigma^3}$
- [Le Cam 1960]: $d_{\text{TV}} \left(\sum_i X_i, \text{Poisson}(\mu) \right) \leq \sum_i p_i^2$
- [Chen-Goldstein-Shao 2011]: $d_{\text{TV}} \left(\sum_i X_i, Z(\mu, \sigma^2) \right) \leq \frac{O(1)}{\sigma}$
- [Röllin 2007]: $d_{\text{TV}} \left(\sum_i X_i, TP(\mu, \sigma^2) \right) \leq \frac{1}{\sigma} + \frac{2}{\sigma^2}$

$\text{Poisson}(\sigma^2 + \{\mu - \sigma^2\}) + \lfloor \mu - \sigma^2 \rfloor$

TP stands for “translated Poisson”

Structure of PBDs

- [Berry 1941, Esseen 1942]: $d_K \left(\sum_i X_i, N(\mu, \sigma^2) \right) \leq C \frac{\mu}{\sigma^3}$
- [Le Cam 1960]: $d_{TV} \left(\sum_i X_i, \text{Poisson}(\mu) \right) \leq \sum_i p_i^2$
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- [Röllin 2007]: $d_{TV} \left(\sum_i X_i, TP(\mu, \sigma^2) \right) \leq \frac{1}{\sigma} + \frac{2}{\sigma^2}$
- ...
- Bounds *only use first two moments*
 - Question 1: Bounds for arbitrary approximation accuracy ε ?
 - Question 2: Distance of two PBDs with same first two moments?
- Approximating distributions *are from a different family*
 - Question 3: Are there meaningful *proper* approximations?

The first $\log(1/\varepsilon)$ -moments suffice

[D-Papadimitriou '09]: Let $X = \sum_i X_i$ and $Y = \sum_i Y_i$ be two PBDs s.t. $\mathbb{E}[X_i] \leq 1/2$ and $\mathbb{E}[Y_i] \leq 1/2$ for all i .

If $\mathbb{E}[X^\ell] = \mathbb{E}[Y^\ell], \forall \ell = 1, \dots, d$

then: $d_{\text{TV}}(X, Y) \leq 2^{-\Omega(d)}$.

Corollary: For all $\varepsilon > 0$, agreement in the first $\log(1/\varepsilon)$ moments suffices for variation distance ε .

The Structure of PBDs

- S_n : set of all PBDs on n variables
- **[D-Papadimitriou '09]**: For all $\epsilon > 0$, there exists a *proper* ϵ -cover $S_{n,\epsilon} \subseteq S_n$ of size:

$$|S_{n,\epsilon}| \leq n^2 + n \cdot (1/\epsilon)^{O(\log^2(1/\epsilon))}$$

\propto : integer multiple

$$\forall P \in S_n, \exists Q \in S_{n,\epsilon} \text{ s.t. } d_{\text{TV}}(P, Q) \leq \epsilon$$

Naïve upper bound for cover size: $|S_{n,\epsilon}| \leq \left(\frac{n}{\epsilon}\right)^n$

- obtained by discretizing every X_i so that its expectation is $\propto \frac{\epsilon}{n}$
which suffices given that:

$$d_{\text{TV}}(\sum_i X_i, \sum_i Y_i) \leq \sum_i d_{\text{TV}}(X_i, Y_i) = \sum_i |\mathbb{E}[X_i] - \mathbb{E}[Y_i]|$$

The Structure of PBDs

- S_n : set of all PBDs on n variables
- **[D-Papadimitriou '09]**: For all $\epsilon > 0$, there exists a *proper* ϵ -cover $S_{n,\epsilon} \subseteq S_n$ of size:

$$|S_{n,\epsilon}| \leq \underbrace{n^2}_{\text{Binomials}} + \underbrace{n \cdot (1/\epsilon)^{O(\log^2(1/\epsilon))}}_{\text{shifted-sparse PBDs}}$$

- In particular:

$$S_{n,\epsilon} = \left[\begin{array}{c} \text{Binomials} \\ \text{Bin}(n', p) \\ n' \leq n \text{ and } p \propto \frac{1}{n} \end{array} \right] \cup$$

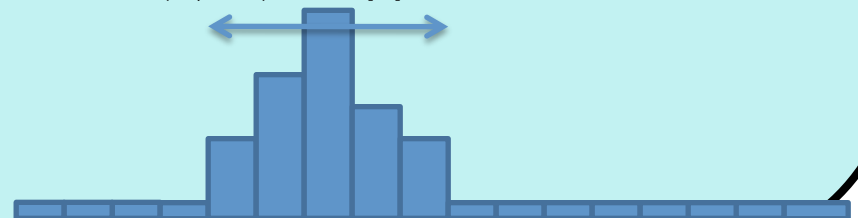
2-parameter distributions

only keep subset of these with different $\log(1/\epsilon)$ first moments

shifted-sparse PBDs

$$\left[\begin{array}{c} n' + \sum_{i=1}^{1/\epsilon^3} Y_i \\ n' \leq n \text{ and } \mathbb{E}[Y_i] \propto \epsilon^2, \forall i \end{array} \right]$$

$O(1/\epsilon^3)$ -support



The Structure of PBDs

- S_n : set of all PBDs on n variables
- **[D-Papadimitriou '09]**: For all $\varepsilon > 0$, there exists a **proper** ε -cover $S_{n,\varepsilon} \subseteq S_n$ of size:

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- In particular:

$$S_{n,\varepsilon} = \left[\begin{array}{c} \text{Binomials} \\ \text{Bin}(n', p) \\ n' \leq n \text{ and } p \propto \frac{1}{n} \end{array} \right] \cup \left[\begin{array}{c} \text{shifted-sparse PBDs} \\ n' + \sum_{i=1}^{1/\varepsilon^3} Y_i \\ n' \leq n \text{ and } \mathbb{E}[Y_i] \propto \varepsilon^2, \forall i \end{array} \right]$$

- **Corollary:** For all $\varepsilon > 0$, every PBD on n variables is either ε -close to a Binomial or ε -close to a shifted PBD on $1/\varepsilon^3$ variables.

Implications to Learning

[D-Diakonikolas-Servedio'12]: Let \mathbf{P} be an unknown PBD in S_n .

- **[Properly Learning PBDs]** Given $\tilde{O}(1/\epsilon^2)$ independent draws from \mathbf{P} and computation time

$$\left(\frac{1}{\epsilon}\right)^{O(\log^2 1/\epsilon)} \cdot \log n$$

can compute a PBD \mathbf{Q} such that $d_{\text{TV}}(\mathbf{P}, \mathbf{Q}) < \epsilon$.

- Any algorithm requires $\Omega(1/\epsilon^2)$ samples (even for $n=1$).

Proof of Learning Result (Attempt 1)

- Use a cover based approach
- **[D-Kamath'14, Acharya et al'14]:** Suppose F_ϵ is an ϵ -cover (in TV distance) of a family of distributions F .

Then can learn any $P \in F$ to within $O(\epsilon)$ -distance using $O\left(\frac{\log |F_\epsilon|}{\epsilon^2}\right)$ samples from P , in time $O\left(\frac{|F_\epsilon| \log |F_\epsilon|}{\epsilon^2}\right)$.

- **How?** Setup tournament among distributions in F_ϵ whose winner cannot be too far from P .
- *Improves long line of similar algorithms **[Devroyé-Lugosi'01, etc]** quadratically in the runtime by designing a better tournament*

- **In our PBD context:** Exists cover of size $n^2 + n \cdot (1/\epsilon)^{O(\log^2(1/\epsilon))}$

→ sample complexity of problem is $\tilde{O}\left(\frac{\log n}{\epsilon^2}\right)$

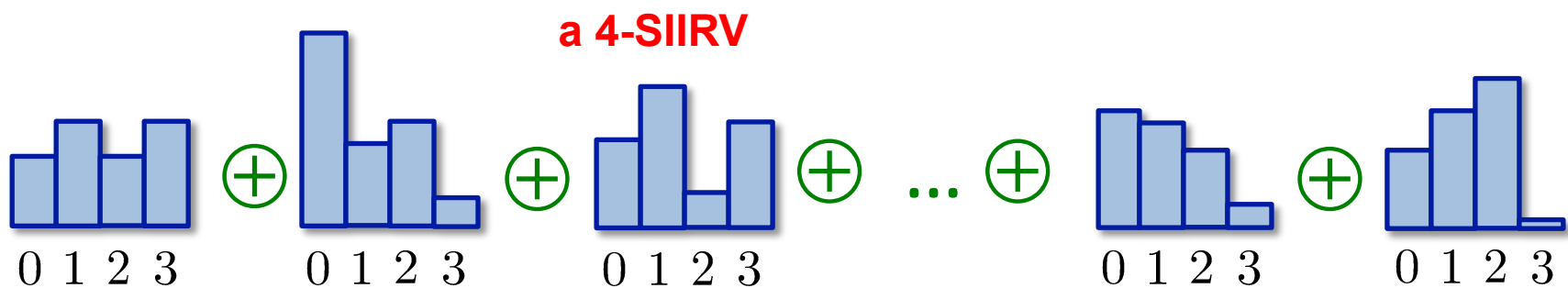
- Fell short from our goal of $\tilde{O}(1/\epsilon^2)$.

Proof of Learning Result (Attempt 2)

- Exploit not just the *size* of the cover, but also its *structure*.
 - We know that every PBD $X = \sum_i X_i$ is ϵ -close to
 - A binomial: $\text{Bin}(n', p)$, $n' \leq n$
 - OR a shifted PBD on $1/\epsilon^2$ variables: $n' + \sum_{i=1}^{1/\epsilon^3} Y_i$, $n' \leq n$
1. Using $O(1/\epsilon^2)$ samples estimate mean and variance of X .
 2. Find Binomial distribution D_1 matching learned mean and variance.
3. In this case, all but ϵ probability mass of X is on support of length $1/\epsilon^3$.
 4. With $\tilde{O}(1/\epsilon^2)$ samples:
 - i. Find the support of $1-\epsilon$ mass of X . This gives estimate of shift n' .
 - ii. Run tournament on ϵ -subcover of shifted by $\approx n'$ PBDs on $1/\epsilon^3$ -variables
 - iii. Let D_2 be the winner of the tournament.
5. Run tournament between D_1 and D_2 .

Menu

- *Refresher*
- *Objectives for this talk*
- *PBD Structure and Learning*
- ***k*-SIIRV Structure and Learning**

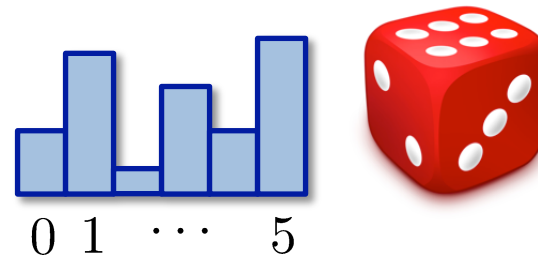


Terminology

k -IRV: Integer-valued Random Variable supported on $\{0, 1, \dots, k-1\}$

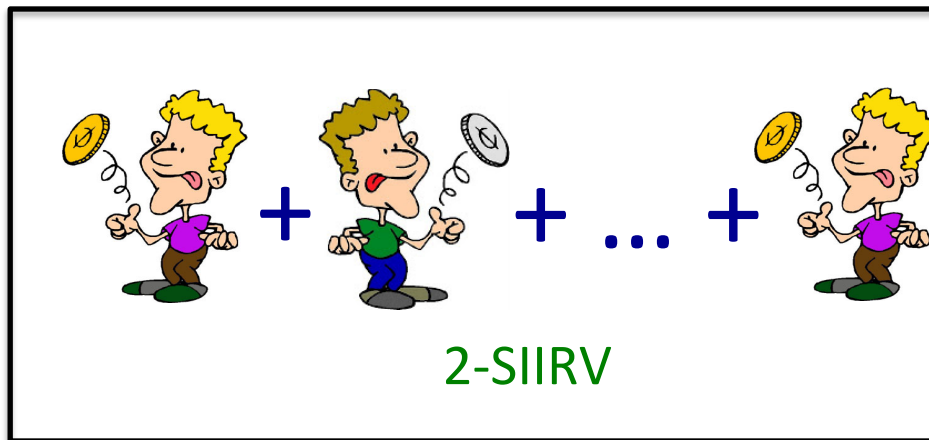


2-IRV



6-IRV

k -SIIRV: Sum of n Independent (*not necessarily identical*) k -IRVs



2-SIIRV



k -SIIRV

Structure ✓ Learning from $\tilde{O}(1/\epsilon^2)$ samples ✓



from 2 to k -SIIRVs: a whole new ball game

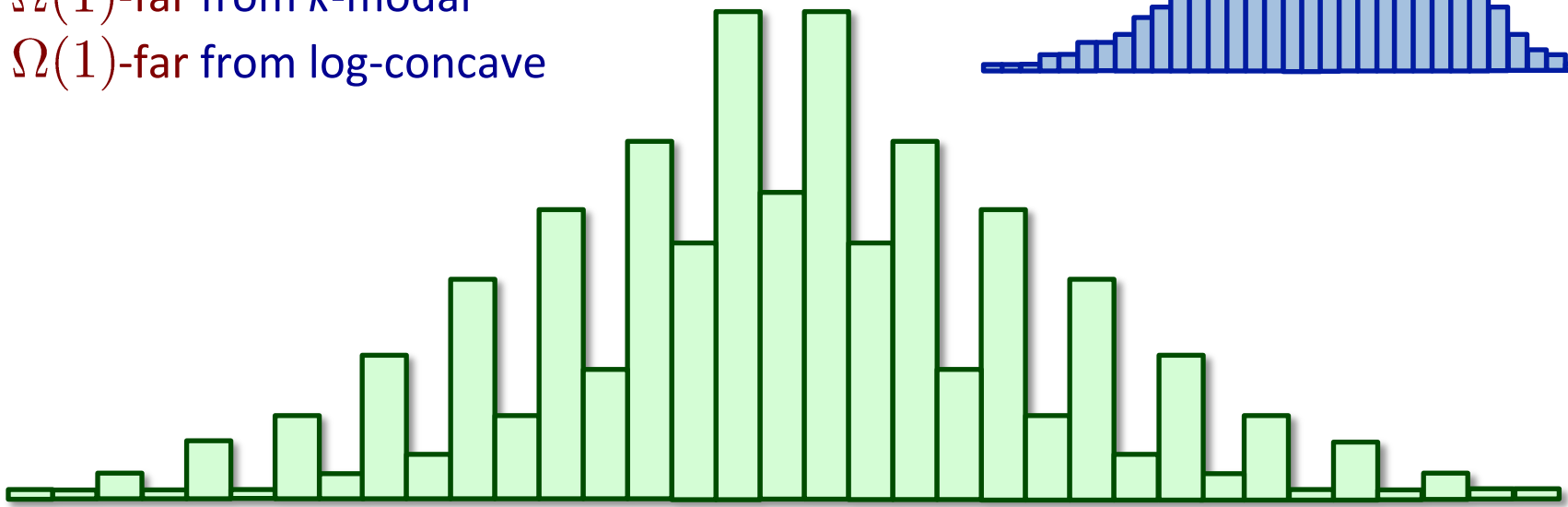
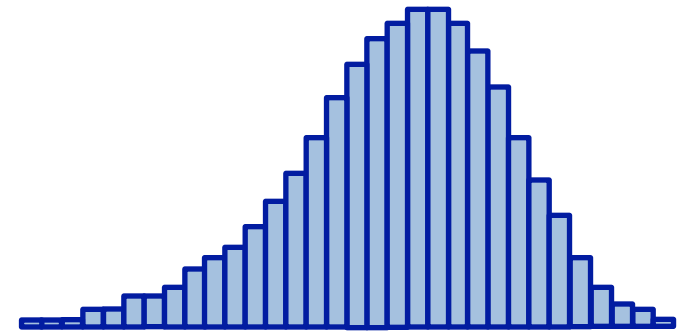
Even just 3-SIIRVs have significantly richer structure than 2-SIIRVs

2-SIIRVs : unimodal, log-concave

3-SIIRVs :

$\Omega(1)$ -far from k -modal

$\Omega(1)$ -far from log-concave

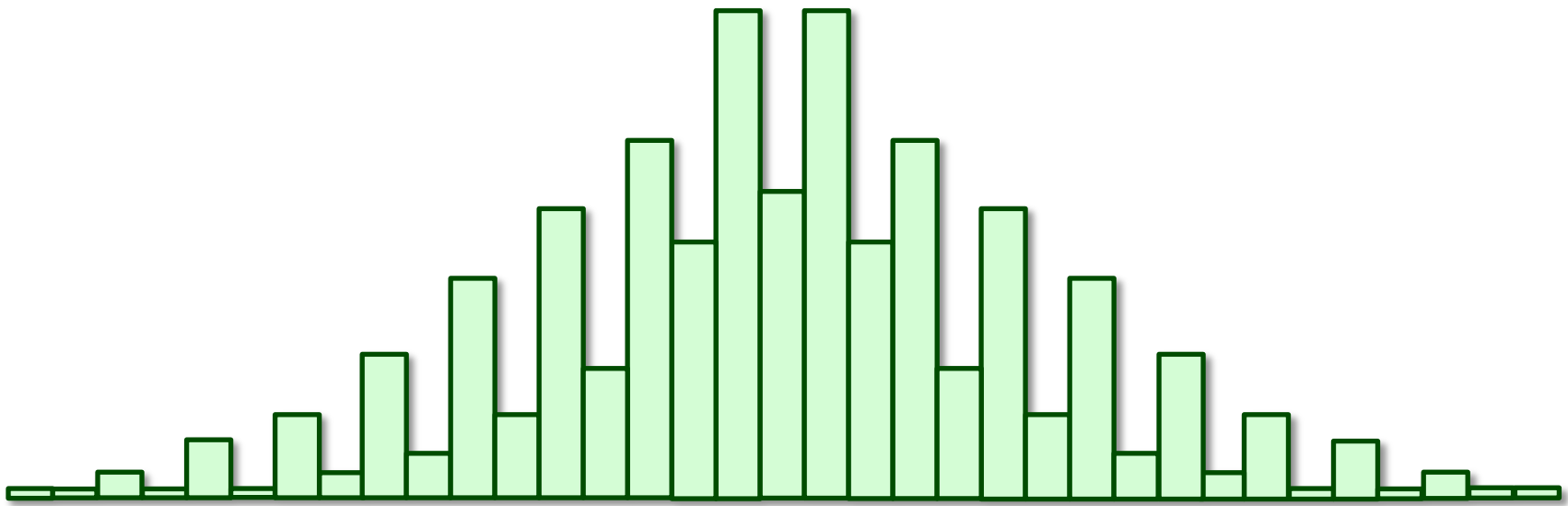


$Y + 2 \cdot Z$, where: $Y \sim \text{Bernoulli}(1/3)$ and $Z \sim \text{Bin}(64, 1/2)$

Structure of k -SIIRVs

- [Berry 1941, Esseen 1942]: If X_1, \dots, X_n are independent and bounded then

$$d_K \left(\sum_i X_i, \mathcal{N}(\mu, \sigma^2) \right) \leq C \frac{\sum_i \mathbb{E}[|X_i|^3]}{\sigma^3}$$



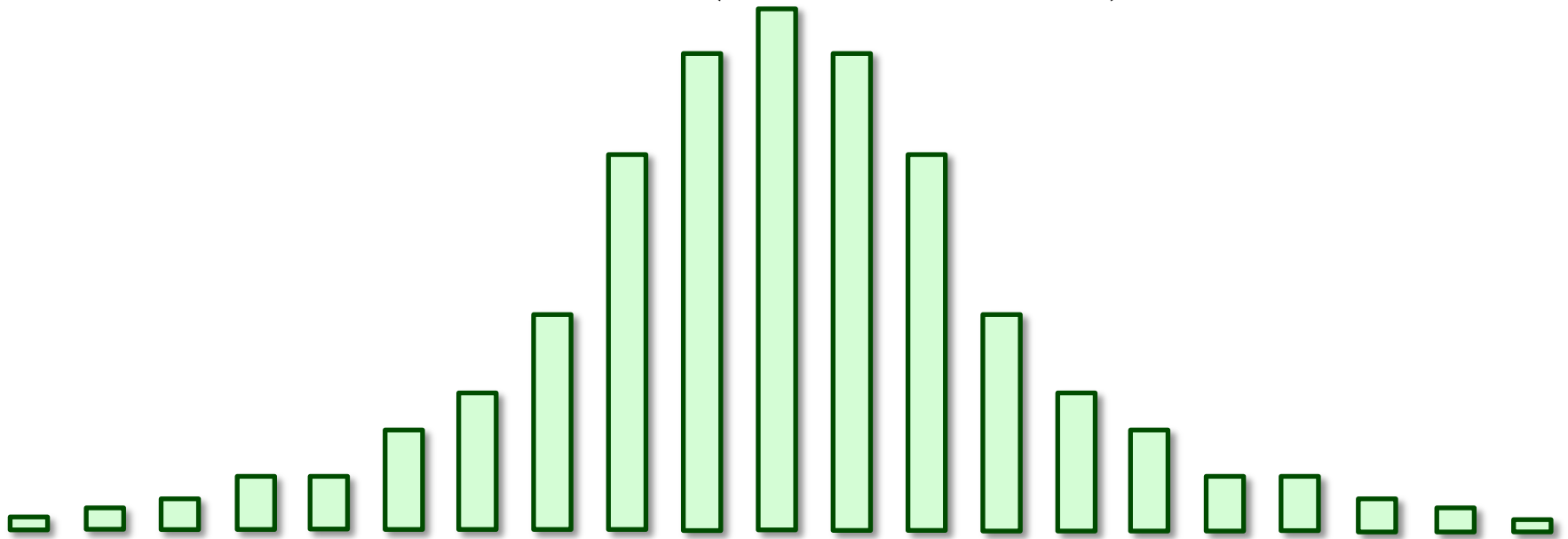
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- Clearly, in general: $d_{\text{TV}} \left(\sum_i X_i, Z(\mu, \sigma^2) \right) = \Omega(1)$
- Conditions under which $d_{\text{TV}} \left(\sum_i X_i, Z(\mu, \sigma^2) \right) = o(1)$?



$2 \cdot Z$, where: $Z \sim \text{Bin}(n, 1/2)$

Structure of k -SIIRVs

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- Conditions under which $d_{\text{TV}} \left(\sum_i X_i, Z(\mu, \sigma^2) \right) = o(1)$?
- [Chen-Goldstein-Shao 2011]: If X_1, \dots, X_n are independent k -IRVs and

$$d_{\text{TV}} \left(\sum_{j \neq i} X_j, \sum_{j \neq i} X_j + 1 \right) \leq \delta, \forall i$$

then

$$d_{\text{TV}} \left(\sum_i X_i, Z(\mu, \sigma^2) \right) = O(k) \left(\frac{1}{\sigma} + \delta \right)$$

[Daskalakis-Diakonikolas-O'Donnell-Servedio-Tan]

[DDOST'13]: Let X be a k -SIIRV with $\text{Var}[X] \geq \text{poly}(k/\epsilon)$.

Then X is ϵ -close to $cZ + Y$, where

- $c \in \{1, 2, \dots, k-1\}$
 - Z = discretized normal
 - Y = c -IRV
- Y, Z : independent

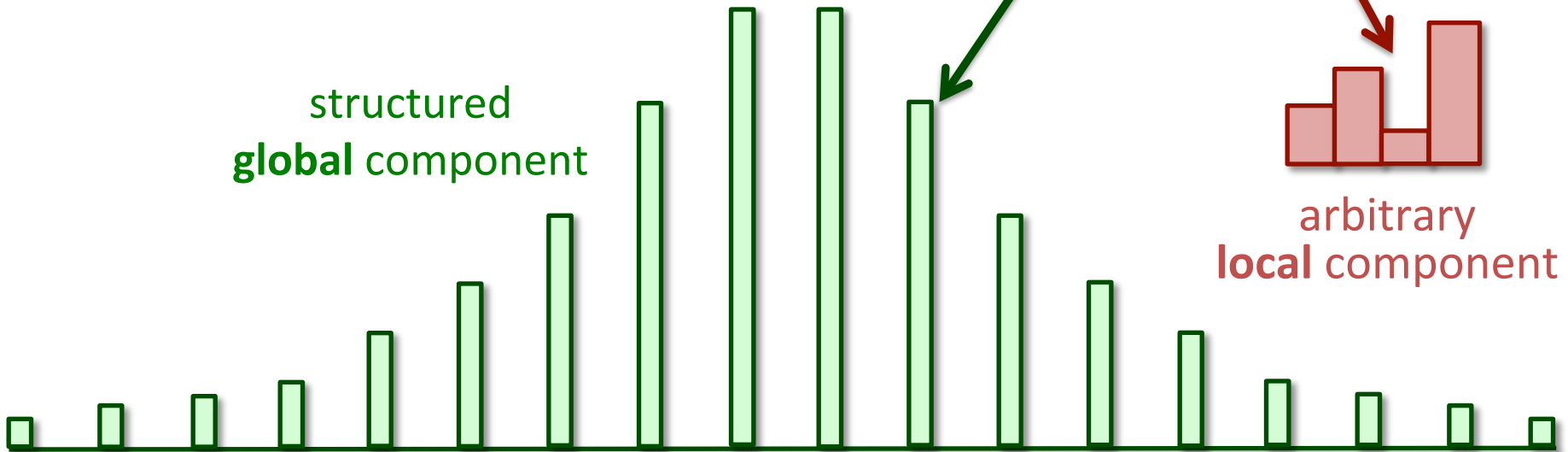


X

$$\approx cZ + Y$$

structured
global component

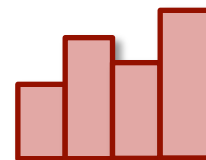
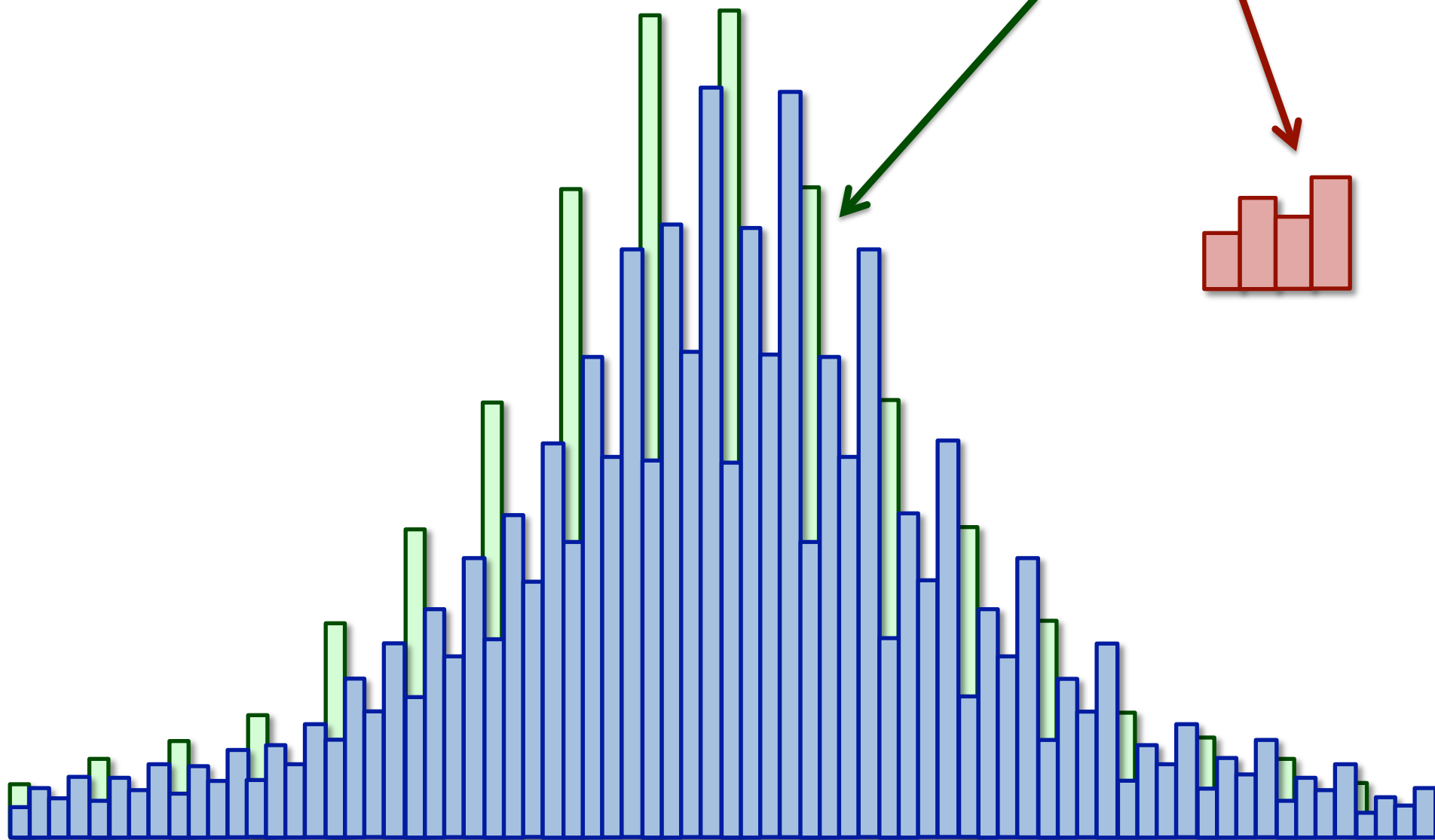
arbitrary
local component





X

$$\approx cZ + Y$$



Structure of k -SIIRVs

[DDOST'13]: Let X be a k -SIIRV with $\text{Var}[X] \geq \text{poly}(k/\epsilon)$.

Then X is ϵ -close to $cZ + Y$, where

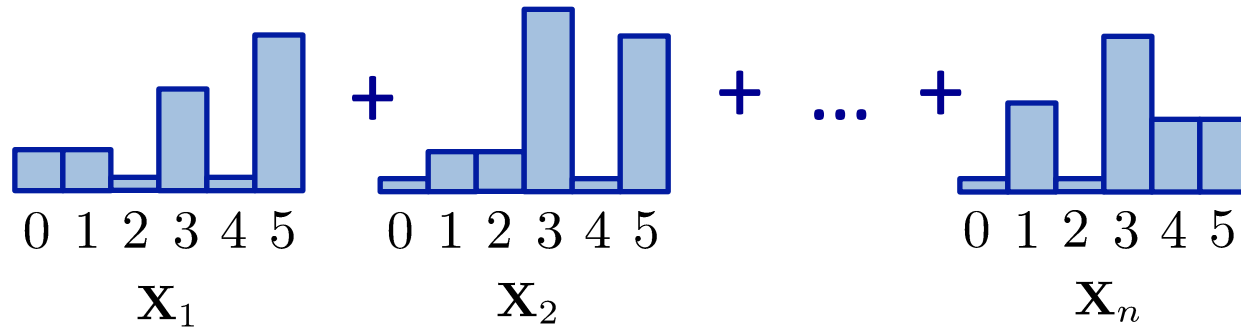
- $c \in \{1, 2, \dots, k-1\}$
- $Z = \text{discretized normal}$ Y, Z : independent
- $Y = c\text{-IRV}$

Corollary: Let X be an arbitrary k -SIIRV. For all $\epsilon > 0$, X is ϵ -close to:

- a $\text{poly}(k/\epsilon)$ -IRV
- *OR* $cZ + Y$, where:
 - $c \in \{1, \dots, k-1\}$
 - $Z = \text{discretized normal}$
 - $Y = c\text{-IRV}$

Proof of Structural Theorem

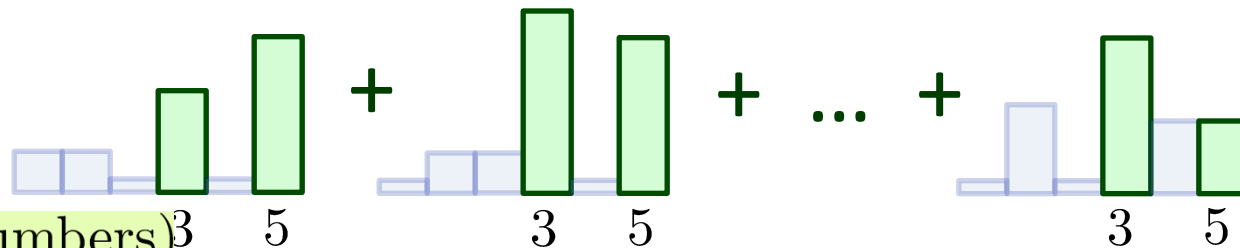
X



cZ

structured **global**
component

Heavy numbers: $\sum_{i=1}^n \Pr[X_i = b]$ large

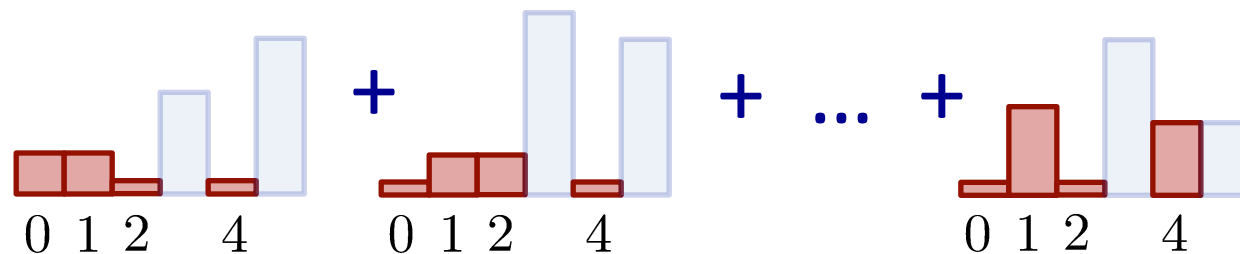


$c = \gcd(\text{heavy numbers})$

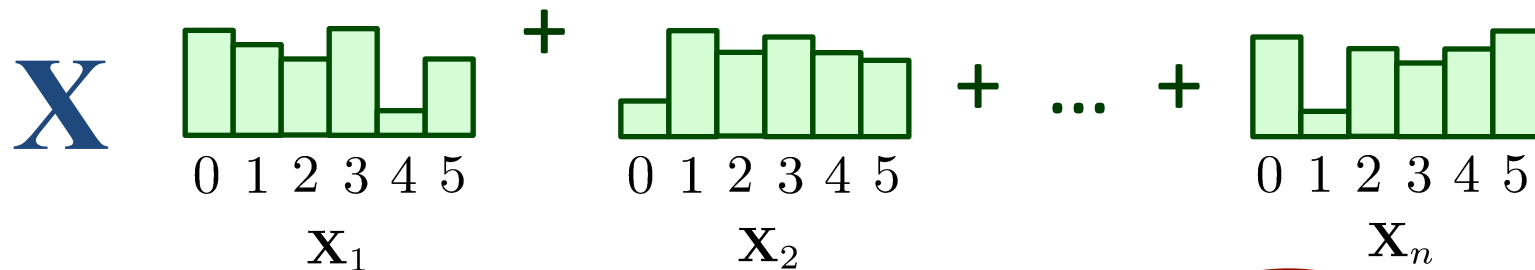
Y

arbitrary **local**
component

Light numbers: $\sum_{i=1}^n \Pr[X_i = b]$ small



Special case: *all* numbers heavy



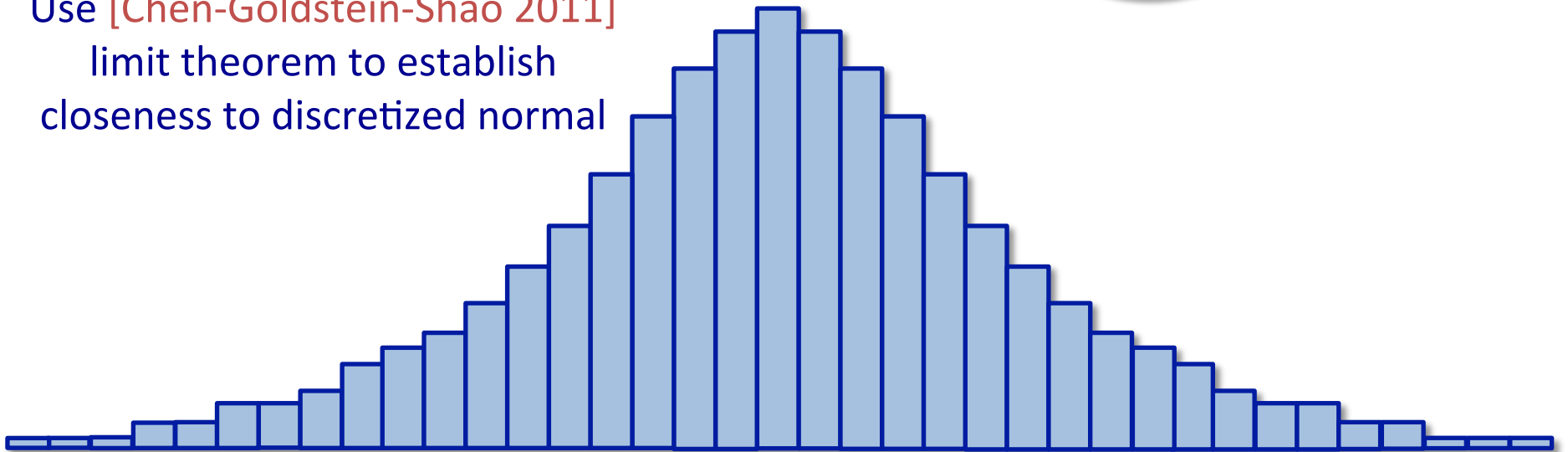
Intuition: No “mod structure” in X

e.g. X equally likely to be 0 or 1 mod 2



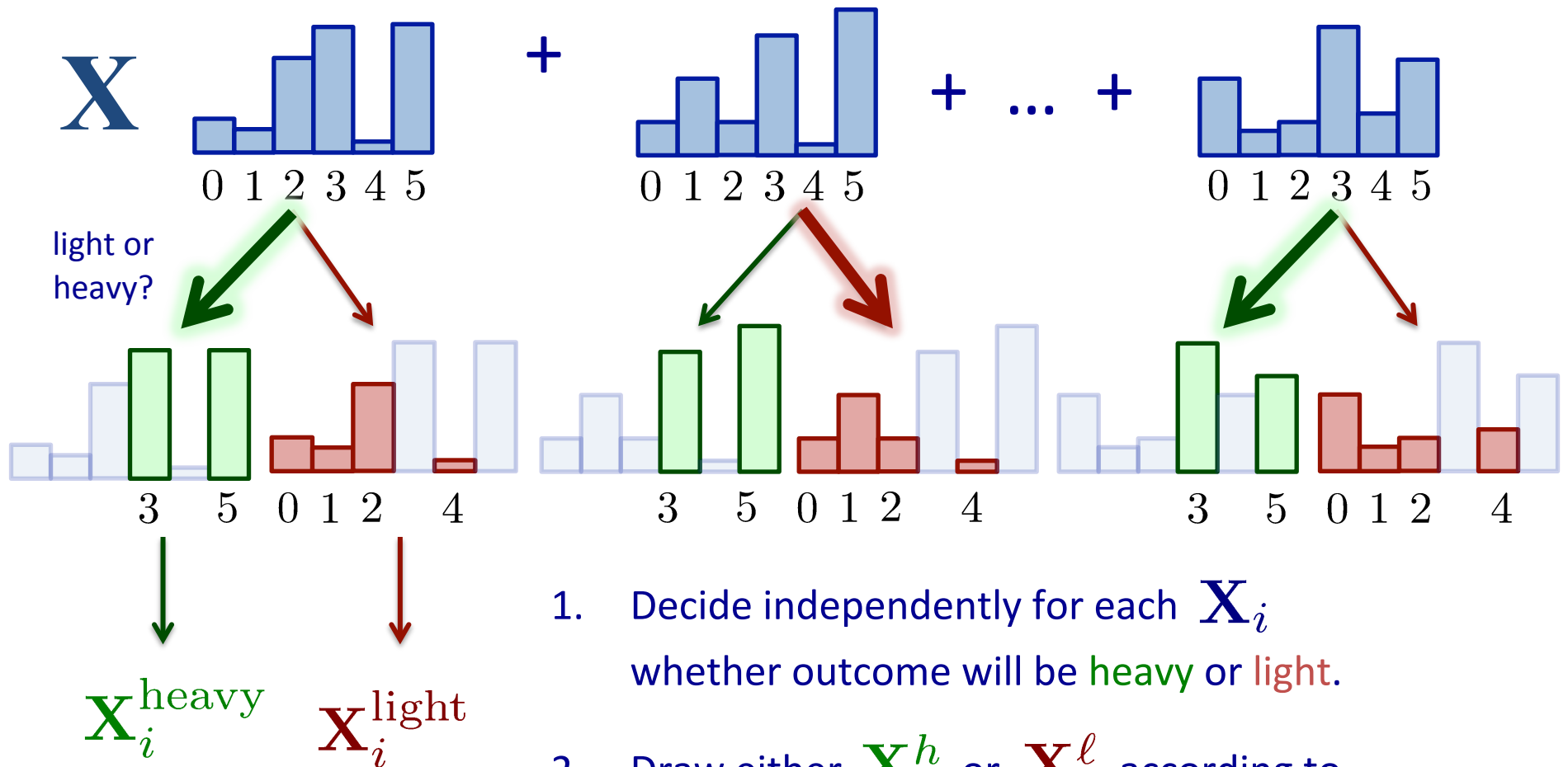
Use [Chen-Goldstein-Shao 2011]

limit theorem to establish
closeness to discretized normal



General Case: indirect sampling procedure

$\{3, 5\}$ heavy, $\{0, 1, 2, 4\}$ light



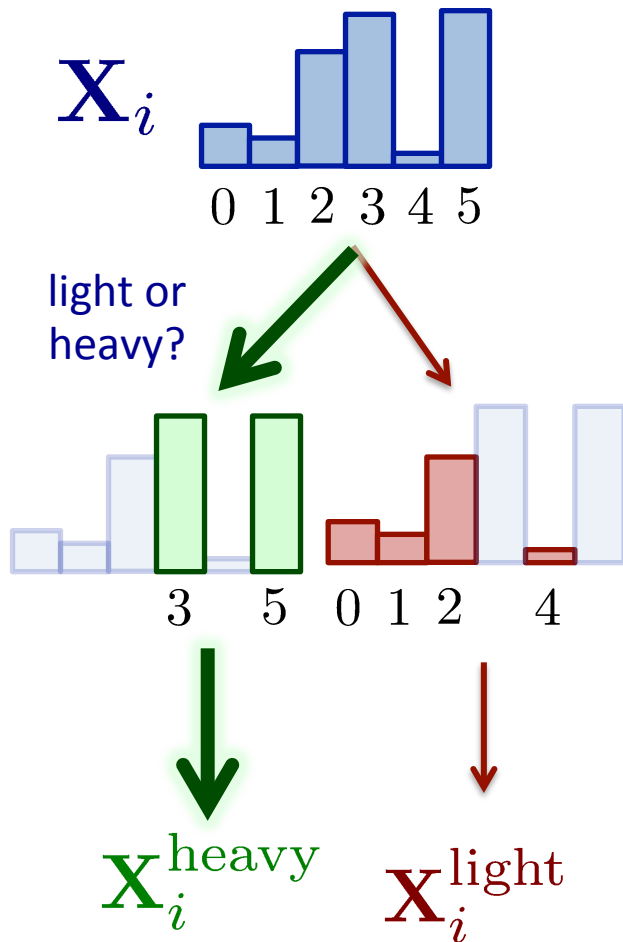
1. Decide independently for each X_i whether outcome will be heavy or light.
2. Draw either X_i^h or X_i^l according to respective conditional distributions.

Analysis

Every outcome O of Stage 1 induces distribution

$$S_O = \sum_{i \in \text{heavy}(\mathcal{O})} \mathbf{x}_i^h + \sum_{j \in \text{light}(\mathcal{O})} \mathbf{x}_j^l$$

\mathbf{S} = mixture of 2^n many S_O 's



Key technical lemma:

With high probability over outcomes O

$$\sum_{i \in \text{heavy}(\mathcal{O})} \mathbf{x}_i^h \approx c \mathbf{Z}$$

where \mathbf{Z} = disc. norm. *independent* of O .

- Proof uses “all numbers heavy” special case
- $c = \text{gcd}(\text{heavy numbers})$

Structure of k -SIIRVs

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Then X is ϵ -close to $cZ + Y$, where

- $c \in \{1, 2, \dots, k-1\}$
- $Z = \text{discretized normal}$ Y, Z : independent
- $Y = c\text{-IRV}$

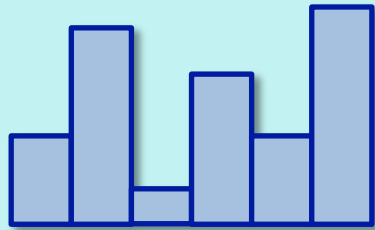
Corollary: Let X be an arbitrary k -SIIRV. For all $\epsilon > 0$, X is ϵ -close to:

- a $\text{poly}(k/\epsilon)$ -IRV
- *OR* $cZ + Y$, where:
 - $c \in \{1, \dots, k-1\}$
 - $Z = \text{discretized normal}$
 - $Y = c\text{-IRV}$

Learning k -SIIRVs

[DDOST'13]: Let $\mathcal{S}_{n,k}$ be the class of k -SIIRVs, *i.e.* all distributions of a sum $X = \sum_{i=1}^n X_i$ of n independent k -IRVs.

There is an algorithm that learns an arbitrary $\mathbf{P} \in \mathcal{S}_{n,k}$ with time and sample complexity $\text{poly}(k/\epsilon)$, independent of n .



Recall: $\Omega(k/\epsilon^2)$ samples
necessary even for a single k -IRV

Proof of Learning Result for k -SIIRVs

Corollary: Let X be an arbitrary k -SIIRV. For all $\epsilon > 0$, X is ϵ -close to:

- a $\text{poly}(k/\epsilon)$ -IRV
- *OR* $cZ + Y$, where:
 - $c \in \{1, \dots, k-1\}$
 - $Z = \text{discretized normal}$
 - $Y = c\text{-IRV}$

- If X is ϵ -close to $\text{poly}(k/\epsilon)$ -IRV: easy to learn from $\text{poly}(k/\epsilon)$ samples
- Else: guess $c \in \{1, \dots, k-1\}$.
- Learn Z from conditional distribution on integers: $0 \bmod k$
- Learn Y as the appropriate mixing distribution
- Run tournament to choose among $(k+1)$ distributions generated.

Summary

1. Ilias discussed how **shape** restrictions on a distribution (monotonicity, k -modality, log-concavity) permit faster learning algorithms
2. I discussed how **syntactic** restrictions on a distribution permit even faster learning:
 - PBDs on n variables have support $\{0, \dots, n\}$ but can be learned from $\tilde{O}(1/\epsilon^2)$ samples
 - k -SIIRVS on n variables have support $\{0, \dots, n(k-1)\}$ but can be learned from $\text{poly}(k/\epsilon)$ samples
3. In turn, finding these improved algorithms requires stronger limit theorems, and tighter structural results for these distributions.
4. **Take-away 1:** Every PBD on n variables is ϵ -close to a Binomial or ϵ -close to a shifted PBD on $1/\epsilon^3$ variables.

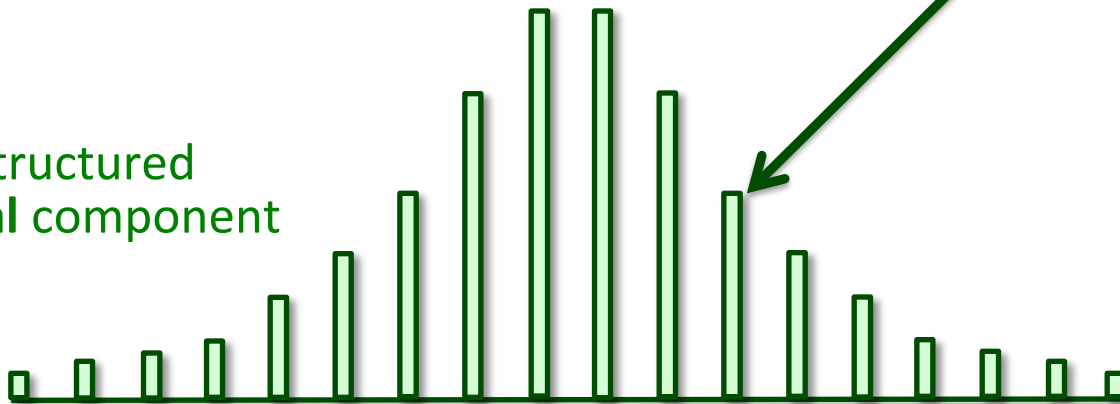
Summary

5. **Take-away 2:** Every k -SIIRV X on n variables is ε -close to a distribution of support $\text{poly}(k/\varepsilon)$ OR

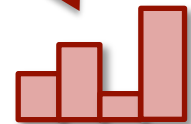


$$\approx_{\varepsilon} cZ + Y$$

structured
global component



arbitrary
local component



Future Directions

6. Super Concrete Open Problems:

- Properly Learn PBDs from $1/\varepsilon^2$ samples in polynomial time (our algorithm was quasi-polynomial)
- Maybe further improve PBD cover size to $\text{poly}(n/\varepsilon)$
- Properly learn k -SIIRVs (our learner outputs a $\text{poly}(k/\varepsilon)$ -SIIRV)
- Proper covers of k -SIIRVs

7. Multi-dimensional Case:

- Learn d -dimensional k -flat distributions, in $\text{poly}(d k/\varepsilon)$ time.
Possible from $\text{poly}(d k/\varepsilon)$ samples (information theoretically)
- Generalize structural/learning results to Poisson Multinomial Distributions
Preliminary results with Kamath and Tzamos

Thanks!