High-Dimensional Robust Mean Estimation in Nearly-Linear Time







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Mean Estimation

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- *Goal:* Learn μ^* .

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• Empirical mean
$$\widehat{\mu} = \frac{1}{N} \sum_{i=1}^{N} X_i$$
 works:
 $\|\widehat{\mu} - \mu^*\|_2 \le \epsilon$ when $N = \Omega(d/\epsilon^2)$.

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Robust Mean Estimation

- *Input:* an ϵ -corrupted set of N samples $\{X_1, \ldots, X_N\}$ drawn from an unknown distribution D on \mathbb{R}^d with mean μ^* .
- *Goal:* Learn μ^* in ℓ_2 -norm.

Robustly learn μ^* given ϵ -corrupted samples from $\mathcal{N}(\mu^*, I)$:

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Geometric Median	$O(\epsilon \sqrt{d})$	Yes
Tournament	$O(\epsilon)$	No
Pruning	$O(\epsilon \sqrt{d})$	Yes
RANSAC	∞	Yes

Robustly learn μ^{\star} given $\epsilon\text{-corrupted}$ samples from $\mathcal{N}(\mu^{\star},I)$:

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[LRV'16]	$O(\epsilon \sqrt{\log d})$	Yes
[DKKLMS'16]	$O(\epsilon \sqrt{\log(1/\epsilon)})$	Yes

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This paper	$O(\epsilon \sqrt{\log(1/\epsilon)})$	$\widetilde{O}(Nd/\epsilon^6)$

All these algorithms have sample complexity $N = O(d/\delta^2)$.

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Our Results

Robustly learn μ^* given ϵ -corrupted samples from D on \mathbb{R}^d .

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Sub-Gaussian	$O(\epsilon \sqrt{\log(1/\epsilon)})$	$O(d/\delta^2)$	$\widetilde{O}(Nd/c^6)$
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Our result provides a faster implementation of such a subroutine, hence yields faster robust algorithms for all these problems.

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Good Weights

minimize
$$\lambda_{\max} \left(\sum_{i=1}^{N} w_i (X_i - \mu^*) (X_i - \mu^*)^{\mathsf{T}} \right)$$

subject to $w \in \Delta_{N,\epsilon} \left(\sum_i w_i = 1 \text{ and } 0 \le w_i \le \frac{1}{(1-\epsilon)N} \right)$

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Lemma ([DKKLMS'16])

If we can find a near-optimal solution w, we can output $\widehat{\mu}_w = \sum_i w_i X_i$.

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Lemma ([DKKLMS'16])

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Our Approach

Idea: guess the mean ν and solve the SDP with parameter ν .

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Primal SDP (with parameter ν) minimize $\lambda_{\max} \left(\sum_{i=1}^{N} w_i (X_i - \nu) (X_i - \nu)^{\top} \right)$ subject to $w \in \Delta_{N, \epsilon}$

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- a near-optimal solution w to the primal SDP give a good answer $\widehat{\mu}_w$, or
- a near-optimal solution to the dual SDP yields a new guess ν' that is closer to μ^* by a constant factor.

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Iteratively move ν closer to μ^* using the dual SDP, until primal SDP has a good solution and we can output $\widehat{\mu}_w$.

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Dual SDP

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maximize Mean of the smallest $(1 - \epsilon)$ -fraction of $((X_i - \nu)^{\top} M(X_i - \nu))_{i=1}^N$ subject to $M \ge 0, \operatorname{tr}(M) \le 1$

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- If the solution is rank-one: $M = yy^{T}$, then in the direction of *y*, the variance is large no matter how we reweight the samples.
- Intuition: When ν is far from μ^* , y should align with $(\nu \mu^*)$.

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Conditions on the Good Samples

We require the following deterministic conditions on the good samples:

Concentration Bounds (for Sub-Gaussian Distributions)

For all
$$w \in \Delta_{N,\epsilon}$$
 (G is the set of good samples):

$$\left\| \sum_{i \in G} w_i (X_i - \mu^*) \right\|_2 \le O(\epsilon \sqrt{\log(1/\epsilon)}) =: \delta_1,$$

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This is the only place we use the sub-Gaussian assumption.

Lemma

When $\|\nu - \mu^*\|_2 \ge \Omega(\beta)$, $1 + 0.99 \|\nu - \mu^*\|_2^2 \le OPT_{\nu} \le 1 + 1.01 \|\nu - \mu^*\|_2^2$.

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Proof

$$OPT_{\nu} \leq \lambda_{\max}\left(\sum_{i=1}^{N} w_i (X_i - \nu) (X_i - \nu)^{\mathsf{T}}\right) = \max_{\|y\|_2 = 1} \sum_{i \in G} w_i (X_i - \nu, y)^2$$

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$$= \max_{\|y\|_2 = 1} \left(\sum_{i \in G} w_i \langle X_i - \mu^*, y \rangle^2 + \langle \mu^* - \nu, y \rangle^2 + 2 \langle \sum_{i \in G} w_i (X_i - \mu^*), y \rangle \langle \mu^* - \nu, y \rangle \right)$$

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$$\begin{aligned} \text{OPT}_{\nu} &\leq \lambda_{\max} \left(\sum_{i=1}^{N} w_i (X_i - \nu) (X_i - \nu)^{\mathsf{T}} \right) = \max_{\|y\|_2 = 1} \sum_{i \in G} w_i \langle X_i - \nu, y \rangle^2 \\ &= \max_{\|y\|_2 = 1} \left(\sum_{i \in G} w_i \langle X_i - \mu^*, y \rangle^2 + \langle \mu^* - \nu, y \rangle^2 + 2 \langle \sum_{i \in G} w_i (X_i - \mu^*), y \rangle \langle \mu^* - \nu, y \rangle \right) \\ &\leq \max_{\|y\|_2 = 1} \left((1 + \delta_2) + \langle \mu^* - \nu, y \rangle^2 + 2 \delta_1 \langle \mu^* - \nu, y \rangle \right) \end{aligned}$$

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$$OPT_{\nu} \leq \lambda_{\max} \left(\sum_{i=1}^{N} w_i (X_i - \nu) (X_i - \nu)^{\top} \right) = \max_{\|y\|_2 = 1} \sum_{i \in G} w_i \langle X_i - \nu, y \rangle^2$$
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=
$$(1 + \delta_2) + \|\mu^{\star} - \nu\|_2^2 + 2\delta_1 \|\mu^{\star} - \nu\|_2 \qquad (\text{so } \beta = \sqrt{\delta_2} = \sqrt{\epsilon \ln(1/\epsilon)}.)$$

For all $w \in \Delta_{N,2\epsilon}$, if $\|\widehat{\mu}_w - \mu^*\|_2 \ge \Omega(\delta)$, then for all $\nu \in \mathbb{R}^d$,

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Proof sketch:

• ν must be close to μ^* , otherwise $OPT_{\nu} \approx 1 + \|\nu - \mu^*\|_2^2$ is already large.

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Proof sketch:

- ν must be close to μ^* , otherwise $OPT_{\nu} \approx 1 + \|\nu \mu^*\|_2^2$ is already large.
- When ν is close to μ^* , $(X_i \nu)(X_i \nu)^{\top}$ is close to $(X_i \mu^*)(X_i \mu^*)^{\top}$.

Fix an approximately optimal solution M to the dual SDP with parameter ν . If the objective value of M is at least $1 + \Omega(\beta^2)$, then we can find $\nu' \in \mathbb{R}^d$ such that $\|\nu' - \mu^*\|_2 \leq \frac{9}{10} \|\nu - \mu^*\|_2$.

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Intuitively, if the dual SDP throws away all the bad samples, $1 + \|\nu - \mu^*\|_2^2 \approx \text{OPT} \approx \mathbb{E}_{X \in G}[(X - \nu)^\top M(X - \nu)] = \langle M, I + (\nu - \mu^*)(\nu - \mu^*)^\top \rangle.$

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- The top eigenvector of M tells us which direction ν should move.
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Primal SDP (with parameter ν)

minimize
$$\lambda_{\max} \left(\sum_{i=1}^{N} w_i (X_i - \nu) (X_i - \nu)^{\mathsf{T}} \right)$$
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Solving the SDPs Approximately

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Binary search for ρ and check if max $||w||_1 \ge 1$ ($\rho^* = \frac{1}{OPT_{\nu}}$). Need to handle bi-criteria approximations. Algorithm 1: Robust Mean Estimation for Known Covariance Sub-Gaussian

- Let $\nu = \frac{1}{N} \sum_{i=1}^{N} X_i$ be the empirical mean;
- for i = 1 to $O(\log(d \log N/\epsilon))$ do

Compute either

(*i*) a good solution $w \in \mathbb{R}^N$ for the primal SDP with parameters $(\nu, 2\epsilon)$; or (*ii*) a good solution $M \in \mathbb{R}^{d \times d}$ for the dual SDP with parameters (ν, ϵ) ; **if** the objective value of w in primal $SDP \le 1 + 400\epsilon \ln(1/\epsilon)$ **then** | **return** the weighted empirical mean $\widehat{\mu}_w = \sum_{i=1}^N w_i X_i$; **else**

Move ν closer to μ^* using the top eigenvector of M.

Algorithm 2: Robust Mean Estimation for Known Covariance Sub-Gaussian

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Distribution	Error (δ)	# of Samples (N)	Runtime
Sub-Gaussian	$O(\epsilon \sqrt{\log(1/\epsilon)})$	$O(d/\delta^2)$	$\widetilde{O}(NJ/c^6)$
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We hope our work will serve as a starting point for the design of faster algorithms for high-dimensional robust estimation.

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- Robust covariance estimation in time $\widetilde{O}(Nd)/\epsilon^{O(1)}$?